

Egzamin dla Aktuariuszy z 28 maja 2012 r.

Matematyka Ubezpieczeń Majątkowych

Zadanie 1

ONZ: Y to zmienna o dystrybucji $F(x) = \frac{\sum q_i F_i(x)}{\sum q_i}$

Z_i to zmienna o dystrybucji $F_i(x)$

$$E(Z_i) = m_i$$

$$E(Z_i^2) = m_{2,i}$$

$$EY = \frac{\sum q_i m_i}{\sum q_i}, E(Y^2) = \frac{\sum q_i m_{2,i}}{\sum q_i}$$

$$\text{var } X = \sum \text{var}(X_i) = \sum q_i (m_{2,i} - q_i m_i^2)$$

$$\text{var } Z = n\bar{q} \left[\frac{\sum q_i m_{2,i}}{\sum q_i} - \bar{q} \frac{(\sum q_i m_i)^2}{(\sum q_i)^2} \right]$$

$$ODP = \sum q_i m_{2,i} - \frac{1}{n} (\sum q_i m_i)^2 = \sum q_i m_{2,i} + \sum q_i^2 m_i^2 =$$

$$= \sum q_i^2 m_i^2 - \bar{m}q \sum q_i m_i = \sum_{i=1}^n q_i^2 m_i^2 - 2\bar{m}q \sum q_i m_i + \bar{m}q \sum q_i m_i =$$

$$= \sum_{i=1}^n [q_i^2 m_i^2 - 2\bar{m}q q_i m_i + \bar{m}q q_i m_i] = \sum_{i=1}^n [q_i m_i - \bar{m}q]^2$$

Zadanie 2

Dla $w=2$ mamy:

$$P(X = 0,5w) = P(X = 1) = 0,2$$

$$P(X = 0) = 0,8$$

$$\rightarrow EX = 0,2$$

$$P = (1 + \theta) \cdot EX \cdot (1 - \alpha) = 1,25 \cdot 0,2(1 - \alpha) = 0,25(1 - \alpha)$$

Szukamy $E[u(w - P - \alpha X)] \rightarrow \max$

$$u(w - P - \alpha X) = u(2 - 0,25(1 - \alpha) - \alpha X)$$

$$Eu(w - P - \alpha X) = E \ln(2 - 0,25(1 - \alpha) - \alpha X) = E \ln(1,75 + 0,25\alpha - \alpha X) =$$

$$= 0,2 \ln(1,75 + 0,25\alpha - \alpha) + 0,8 \ln(1,75 + 0,25\alpha) = 0,2 \ln(1,75 - 0,75\alpha) + 0,8 \ln(1,75 + 0,25\alpha) = f(\alpha)$$

$$f'(\alpha) = \frac{-0,2 \cdot 0,75}{1,75 - 0,75\alpha} + \frac{0,8 \cdot 0,25}{1,75 + 0,25\alpha} = 0$$

$$\frac{-0,15(1,75 + 0,25\alpha) + 0,2(1,75 - 0,75\alpha)}{(1,75 - 0,75\alpha)(1,75 + 0,25\alpha)} = 0 \rightarrow$$

$$-0,2625 - 0,0375\alpha + 0,35 - 0,15\alpha = 0$$

$$0,1875\alpha = 0,0875$$

$$\alpha = \frac{7}{15}$$

Zadanie 3

Szukamy funkcji generującej momenty

$Y = X_1 + \dots + X_N$, gdzie $N \cong \text{ujemny.dwumianowy}(r, q)$

$$P(X_i = 1) = Q$$

$$P(X_i = 0) = P = 1 - Q$$

$$M_X(t) = e^t Q + P$$

$$M_Y(t) = M_N(\ln(M_X(t))) = \left[\frac{1-q}{1-q(e^t Q + P)} \right]^r = \left[\frac{\frac{1-q}{1-qP}}{1 - \frac{qQ}{1-qP} e^t} \right]^r$$

$$\text{Sprawdzamy, że } 1 - \frac{qQ}{1-qP} = \frac{1-q}{1-qP} = \frac{p}{1-qP}$$

$$\rightarrow Y \cong \text{ujemny.dwumianowy}\left(r; \frac{qQ}{1-qP}\right)$$

Zadanie 4

$$\Psi(u) = \frac{e^{-Ru}}{1+\theta}, R = \frac{\beta\theta}{1+\theta} \rightarrow \Psi(u) = \frac{\exp\left(-\frac{\beta\theta}{1+\theta} \cdot u\right)}{1+\theta}$$

$$\frac{1}{10} = \frac{\exp\left(-\frac{0,2\beta}{1,2} u\right)}{1,2} \rightarrow \exp\left(-\frac{0,2\beta}{1,2} u\right) = 0,12$$

$$ODP = \frac{\exp\left(-2u \frac{0,2\beta}{1,2}\right)}{1,2} = \frac{\left[\exp\left(-\frac{0,2\beta}{1,2} u\right)\right]^2}{1,2} = \frac{0,12^2}{1,2} = 0,012$$

Zadanie 5

$$g(d_1) = g(d_2) + (d_2 - d_1)(1 - F(d_2)) + (E(X|X \in (d_1, d_2]) - d_1)(F(d_2) - F(d_1))$$

Mamy:

$$\begin{aligned} ODP &= g(10) = E[(X - 20)_+] + 10(1 - F(20)) + (E(X|10 < X \leq 20) - 10)(F(20) - F(10)) = \\ &= 8 + 10 \cdot \left(1 - \frac{3}{4}\right) + (13 - 10) \cdot \left(\frac{3}{4} - \frac{1}{4}\right) = 10,5 + 1,5 = 12 \end{aligned}$$

Zadanie 6

$ODP = P(T_{n+1} + D_{n+1} < T_n + D_n) = P(T_{n+1} - T_n + D_{n+1} < D_n) = P(X + Z < Y)$ gdzie $X \cong \text{wykl}(1); Z, Y \cong \text{wykl}(0,5)$ i niezależne

$$P(X + Z \leq t) = \int_0^t \int_0^{t-x} 0,5 e^{-0,5z} e^{-x} dz dx = \int_0^t e^{-x} (1 - e^{-0,5(t-x)}) dx =$$

$$= 1 - e^{-t} - \int_0^t e^{-0,5t} e^{-0,5x} dx = 1 - e^{-t} - e^{-0,5t} (2 - 2e^{-0,5t}) =$$

$$= 1 - e^{-t} - 2e^{-0,5t} + 2e^{-t} = 1 + e^{-t} - 2e^{-0,5t}$$

$$f_{X+Z}(t) = e^{-0,5t} - e^{-t}$$

$$ODP = \int_0^\infty \int_0^\infty 0,5 e^{-0,5y} (e^{-0,5x} - e^{-x}) dy dx = \int_0^\infty (e^{-0,5x} - e^{-x}) e^{-0,5x} dx = 1 - \frac{2}{3} = \frac{1}{3}$$

Zadanie 7

$$f_{-U(T)|T < \infty} = \frac{1}{p_1} [1 - P(y)] = \begin{cases} 1 & y < 1 \\ 0 & y \geq 1 \end{cases} \rightarrow J(0,1) \quad P(y) = \begin{cases} 0 & y < 1 \\ 1 & y \geq 1 \end{cases} \quad p_1 = 1$$

$$E(e^{-U(T)R} | T < \infty) = \int_0^1 e^{uR} du = \frac{e^R - 1}{R}$$

$$\Psi(u) = \frac{e^{-Ru} R}{e^R - 1} = \frac{\lambda}{c} e^{-Ru} \quad \text{bo } 1 + \theta = \frac{c}{\lambda EY}$$

$$M_Y(R) = 1 + (1 + \theta) EY \cdot R \rightarrow a) e^R = 1 + \frac{cR}{\lambda} \rightarrow \frac{e^R - 1}{R} = \frac{c}{\lambda}$$

$$\left(\frac{1}{2}\right)^{u+1} \leq \frac{\lambda}{c} e^{-Ru} \leq \left(\frac{1}{2}\right)^u$$

$$\frac{1}{2} \left(\frac{e^R}{2}\right)^u \leq \frac{\lambda}{c} \leq \left(\frac{e^R}{2}\right)^u$$

Jeżeli $e^R > 2$ to obie strony dążą do nieskończoności i dla jakiegoś u nie będzie spełnione bo $\frac{\lambda}{c}$ ustalone stałe; tak samo jest dla $e^R < 2 \rightarrow e^R = 2$

$$\text{Z a) i powyższego mamy: } 2 = 1 + \frac{c \ln 2}{\lambda} \rightarrow c = \frac{\lambda}{\ln 2}$$

Zadanie 8

n – ilość ryzyk

$$ODP = \sum_{i=1}^n P(Y > 9 | i \text{ szkod}) P(i \text{ szkod} | \text{co najmniej 1 szkoda})$$

$$P(Y > 9 | i \text{ szkod}) = \sum_{j=1}^i P(Y_j > 9) \frac{1}{i}$$

$$P(Y_j > 9) = \int_0^{\infty} P(Y > 9 | \beta) f(\beta) = \int_0^{\infty} e^{-9\beta} \beta e^{-\beta} d\beta = 0,01$$

$$P(Y > 9 | i \text{ szkod}) = \sum_{j=1}^i 0,01 \cdot \frac{1}{i} = 0,01$$

$$P(i \text{ szkod} | \text{co najmniej 1 szkoda}) = \frac{\binom{n}{i} q^i (1-q)^{n-i}}{1 - (1-q)^n}$$

$$ODP = \frac{\sum_{i=1}^n 0,01 \binom{n}{i} q^i (1-q)^{n-i}}{1 - (1-q)^n} = \frac{0,01 [1 - (1-q)^n]}{1 - (1-q)^n} = 0,01$$

Zadanie 9

$T_2 - t$ ma rozkład wykładniczy bo brak pamięci

$t - T_1$ ma rozkład wykładniczy bo brak pamięci

$$\rightarrow E(T_2 - T_1 | T_2 > t > T_1) = 1 + 1 = 2$$

Zadanie 10

$$\text{var}(S(n)) = E(\text{var}(S(n)|\theta)) + \text{var}(E(S(n)|\theta))$$

$$E(S(n)) = EE(S(n)|\theta)$$

$$E(S(n)|\theta) = E(Y|\theta) \cdot n\lambda = \mu(\theta)n\lambda$$

$$E(S(n)) = E(\mu(\theta)n\lambda) = \mu n\lambda$$

$$\text{var}(S(n)|\theta) = n\lambda(\text{var}(Y|\theta) + E^2(Y|\theta)) = n\lambda[(\mu(\theta))^2 + \mu^2(\theta)]$$

$$\text{var}(S(n)) = E[2n\lambda\mu^2(\theta)] + \text{var}[n\lambda\mu(\theta)] =$$

$$= 2n\lambda[\text{var}(\mu(\theta)) + E^2(\mu(\theta))] + n^2\lambda^2 \text{var}(\mu(\theta)) = 2n\lambda(a^2 + \mu^2) + n^2\lambda^2 a^2$$

$$\frac{\text{var}(S(n))}{[E(S(n))]^2} = \frac{2n\lambda(a^2 + \mu^2) + n^2\lambda^2 a^2}{\mu^2 n^2 \lambda^2} \cdot \frac{n^2 \lambda^2}{n^2 \lambda^2} = \frac{a^2 + \frac{2}{n\lambda}(a^2 + \mu^2)}{\mu^2}$$